Incorporating Residential Self-Selection Effects in Activity Time-use Behavior: Formulation and Application of a Joint Mixed Multinomial Logit – Multiple Discrete-Continuous Extreme Value Model

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ABSTRACT
This study presents a joint model system of residential location and activity time-use choices that considers a comprehensive set of activity-travel environment (ATE) variables, as well as sociodemographic variables, as determinants of individual weekday activity time-use choices. The model system takes the form of a joint mixed Multinomial Logit–Multiple Discrete-Continuous Extreme Value (MNL–MDCEV) structure that (a) accommodates differential sensitivity to the ATE attributes due to both observed and unobserved individual and household attributes, and (b) controls for the self selection of individuals into neighborhoods based on observed and unobserved activity time-use preferences. The joint model system is estimated on a sample of 2793 households and individuals residing in Alameda County in the San Francisco Bay Area.

The model results indicate the significant presence of residential sorting effects based on time-use preferences. For instance, individuals with bicycles locate themselves into neighborhoods with good bicycling facilities. These same individuals also have a preference for physically active pure recreation pursuits. Ignoring this effect of bicycle ownership would lead to an inflated estimate of the effect of bicycling facility density on time invested in physically active recreation. Similarly, there are unobserved individual factors (such as fitness consciousness) that make individuals locate in areas with good bicycling facilities and also lead to a high preference for physically active recreation. That is, people who are predisposed to physically active lifestyles tend to self-select themselves into zones with very good bicycling facility density for their residence. These results show the danger of ignoring residential sorting effects when estimating the effects of the activity-travel environment on activity-travel choices. Nevertheless, the findings from our study indicate that modifying the activity-travel environment can lead to small changes in individual activity time-use patterns, even after controlling for residential sorting effects.
1. INTRODUCTION
The primary focus of transportation planning, until the past three decades or so, was to meet long-term mobility needs by providing adequate transportation infrastructure supply. In such a mobility-centric, supply-oriented, planning process, the main role of travel demand models was to predict aggregate travel demand for long-term socio-economic scenarios, and for alternative transportation system characteristics and land-use configurations.

Over the past three decades, however, because of rapid economic growth, increasing auto dependency, accelerating urban sprawl, escalating capital costs of new infrastructure, and increasing concerns regarding air-quality deterioration and traffic congestion, the mobility-centric, supply-oriented, focus of transportation planning has expanded to include the objectives of (a) promoting sustainable communities and urban areas by integrating transportation planning with land-use planning, and (b) addressing mobility needs and problems by managing travel demand within the available transportation supply. Consequently, and correspondingly, there has been an increasing interest in (a) land-use policies (such as mixed land-use development and transit oriented development) that attempt to modify the land-use configuration in an effort to reduce auto-oriented travel and promote other means of transportation, and (b) travel demand management strategies, such as congestion pricing, that attempt to change transportation service characteristics to influence individual travel behavior and control aggregate travel demand.

The interest in analyzing the potential of land-use and travel demand management policies to manage travel demand, in turn, has led to a shift in travel demand modeling from the statistical prediction of aggregate-level, long-term, travel demand to understanding disaggregate-level (i.e., individual-level) behavioral responses to land-use and demand management policies. This is evidenced in the evolution of the field along two directions: (a) Integrated land-use travel demand modeling (to better understand land-use policy impacts), and (b) Activity-based travel demand modeling (to better understand demand management policy effects). Each of these two research directions are discussed in turn in the next two sections, while Section 1.3 positions the research in the current study.

1.1 Integrated Land-use Travel Demand Modeling
In the typical approach to travel demand modeling, land-use attributes are considered predetermined and exogenous, and are used as independent variables to explain travel behavior. Such an approach implicitly assumes a one-directional relationship between land-use and travel demand. In the past decade, there has been an increasing amount of research focused on revisiting this simplistic assumption, and considering more complex inter-relationships that may exist between land-use and travel demand. At the center of this debate and investigation is whether any effect of land-use on travel demand is causal or merely associative (or some combination of the two; see Bhat and Guo, 2007, and Cao et al., 2006a). To understand this issue better, consider a land-use policy to improve bicycling facilities, with the objective of reducing automobile dependence and increasing physically active recreational pursuits. To assess the impacts of such a policy, assume that a data collection effort has been undertaken to examine the bicycling levels of individuals in neighborhoods with different levels of existing bicycling facilities. An analysis of this data may find that individuals residing in neighborhoods with good bicycling facilities pursue more bicycling-related activities. The question is whether this relationship implies that building neighborhoods with good bicycling facilities would result in higher bicycling levels in the overall population (i.e., a causal relationship), or whether this relationship is an artifact of individuals who are bicycling-inclined (because of, say, physical...
fitness consciousness) self-selecting themselves to reside in neighborhoods with good bicycling facilities (i.e., an associative relationship). If the latter “residential sorting process” is at work, building neighborhoods with good bicycling facilities would not result in higher bicycling levels in the overall population, but simply lead to an alteration of spatial residence patterns of the population based on physical fitness consciousness.¹

In reality, the nature of the relationship between land-use and travel behavior may be part causal and part associative. Thus, any attempt to examine the land use-travel behavior connection should disentangle the causal and associative elements of the relationship to inform and contribute to the credible assessment of the impact of land-use policies on travel behavior. Several recent studies have alluded to and/or accommodated this residential sorting process in one of several ways. However, all these earlier studies study the land use and travel behavior relationship by directly focusing on specific travel behavior dimensions, such as trip frequency or trip mileage for one or more trip purposes.² Basically, these earlier studies adopt a trip-based approach to analyze the interactions between land use and travel demand, ignoring the conceptual and behavioral limitations of the trip-based approach (see Jones et al., 1990, Axhausen and Garling, 1992, Kurani and Kitamura, 1996, Bhat and Koppelman, 1999, Bhat et al., 2004, and Vovsha and Bradley, 2005 for detailed discussions of the shortcomings of the trip-based method; space constraints do not allow us to discuss these limitations in this paper).

1.2 Activity-based Travel Demand Modeling
The activity-based approach to travel demand analysis overcomes the conceptual and behavioral inadequacy of the trip-based approach by (1) recognizing that travel is derived from a more fundamental need to perform activities distributed in time and space, (2) focusing on entire sequences and patterns of activities and travel over the course of a day rather than individual trips, (3) recognizing the linkages among activity-travel decisions of an individual across different time periods of the day, (4) explicitly modeling the temporal dimension of activity participations and travel, and (5) accommodating space-time interactions in activities and travel. All of these aspects of the activity-based approach may be traced back to a single fundamental difference between the trip-based approach and the activity-based approach, which is in the way time is conceptualized and represented in the two approaches (see Pas, 1996, Pas and Harvey, 1997, and Bhat and Koppelman, 1999). In the trip-based approach, time is reduced to being simply a "cost" of making a trip. The activity-based approach, on the other hand, treats time as an all-encompassing continuous entity within which individuals make activity/travel

¹ As indicated in Bhat and Guo (2007), the assumption here is that there is an adequate supply of neighborhoods to choose from for individuals who are bicycling-oriented. If there is an undersupply of such neighborhoods, then enhancing bicycling facilities in some neighborhoods would increase overall bicycling activity across the population even if the only process at work is residential sorting.

² See Cao et al. (2005, 2006b), Chatman (2005), Kitamura et al. (1997), Schwanen and Mokhtarian (2003), Boarnet and Sarimento (1998), Greenwald and Boarnet (2001), Khattak and Rodriguez (2005), Handy et al. (2006), Handy and Clifton (2001), and Krizek (2000 and 2003) for analyses of residential self selection in models of trip frequency by one or more modes and/or purposes; Schwanen and Mokhtarian (2005b), Khattak and Rodriguez (2005), Bagley and Mokhtarian (2002), Handy et al. (2005), and Krizek (2000, 2003) for analyses of residential self selection in models of travel mileage by one or more modes; Cervero and Duncan (2002), Hammond (2005), Pinjari et al. (2007), Schwanen and Mokhtarian (2005a), Salon (2006), and Zhang (2006) for commute mode choice analyses that consider residential self selection; and Cervero and Duncan (2003), Greenwald (2003), and Salon (2006) for non-commute mode choice analyses that accommodate residential self selection. The reader is also referred to Bhat and Guo (2007) and Cao et al. (2006a) for an overview of studies that address residential self-selection in examining land use variable impacts on travel dimensions.
participation decisions (see Kurani and Lee-Gosselin, 1996). Thus, the central basis of the activity-based approach is that individuals' activity-travel patterns are a result of their time-use decisions. Individuals have 24 hours in a day (or multiples of 24 hours for longer periods of time) and decide how to use that time among activities and travel (and with whom) subject to their sociodemographic, spatial, temporal, transportation system, and other contextual constraints. In the activity-based approach, the impact of land-use and demand management policies on time-use behavior is an important precursor step to assessing the impact of such polices on individual travel behavior. For example, one may analyze whether improving a neighborhood with walkways, bikeways, and recreational parks encourages individuals to invest more time in physically active recreation pursuits in the place of in-home passive recreation (such as watching television or playing computer games). The travel dimensions then can be “derived” from the changes in time-use and activity-scheduling patterns.

To be sure, there have been several studies of individual time-use in the past decade. However, most of these studies have focused on understanding time-use patterns as a function of individual and household sociodemographics (see, for example, Bhat and Misra, 1999, Chen and Mokhtarian, 2006, Gilebe and Koppelman, 2002, Golob and McNally, 1995, Goulias and Henson, 2006, Harvey and Taylor, 2000, Kapur and Bhat, 2007, Kitamura et al., 1996, Kraan, 1996, Levinson, 1999, Lu and Pas, 1997 and 1999, Meloni et al., 2004, Meloni et al., 2007, Yamamoto and Kitamura, 1999, and Ye and Pendyala, 2005). Such time-use studies, while contributing to the literature in important ways, are unable to examine the impact of land-use policies. Some more recent studies by Bhat and colleagues do include the impact of a comprehensive set of land use attributes in their time-use analyses (see, for example, Bhat, 2005, Bhat et al., 2006, Copperman and Bhat, 2007, and Sener and Bhat, 2007). However, these studies are focused on weekend day time-use behavior and not weekday time-use behavior. Besides, an important limitation of all these earlier activity-based time-use studies is that they do not consider residential sorting effects when assessing the impact of land use attributes. Rather, they assume land use as being pre-determined and exogenous.

1.3 The Current Research
The discussion in the previous sections indicates the substantial earlier research on integrated land use-travel demand modeling and activity-based travel demand modeling. While the research on integrated land use-travel demand modeling has been driven by a need to assess land use policies, the emergence of the activity-based approach may be attributed to the need to understand individual-level behavioral responses to demand management policies. Unfortunately, however, these two streams of research have progressed in rather independent directions. In particular, integrated land use-travel demand models have adopted a trip-based approach, while activity-based travel demand models have adopted a “dis-integrated” land use-travel demand approach.

In the current research, we bring the foregoing two streams of research together in the context of a joint model of traffic analysis zone (TAZ)-level residential choice and individual activity time-use behavior. Specifically, we accommodate residential sorting effects due to observed and unobserved individual characteristics in examining the impact of activity-travel environment (ATE) variables on individual time-use in maintenance activity (grocery shopping, household chores, personal care, etc.) and several types of discretionary activity purposes. The

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3 In the rest of this paper, we will use the more general term ATE instead of land-use. The ATE attributes of a traffic analysis zone (TAZ) include the sociodemographic, natural, and built environment characteristics of the TAZ, as
residential choice-activity time use model system in the paper takes the form of a joint mixed multinomial logit – multiple discrete-continuous extreme value (MNL–MDCEV) model. To our knowledge, this is the first instance in the econometric literature of the development of such a model to jointly analyze an unordered discrete variable (residential location choice in the current context) and multiple discrete-continuous variables (activity participations and time-use decisions in multiple activities in the current context).

The remainder of this paper is organized as follows. Section 2 presents the mathematical structure of the joint model and the estimation procedure. Section 3 discusses the data sources, the ATE measures constructed, and the sample used in the analysis. Section 4 focuses on the empirical results. Section 5 demonstrates an application of the model to predict activity time-use changes in response to changes in specific ATE policies. Finally, Section 6 concludes the paper by summarizing important findings and identifying directions for future research.

2. ECONOMETRIC MODELING FRAMEWORK

2.1 Model Structure

Let $q$ ($q = 1, 2, \ldots, Q$) be an index for the decision-maker, $k$ ($k = 1, 2, \ldots, K$) be the index for activity purpose, and $i$ ($i = 1, 2, \ldots, I$) be the index for the spatial unit of residence. Let $T_q$ be the total amount of time available to individual $q$ for participation in maintenance and discretionary activity purposes, and let $t_q = \{t_{q1}, t_{q2}, \ldots, t_{qk}\}$ be the vector of time investments in maintenance activity $(t_{q1})$ and discretionary activities $(t_{q2}, t_{q3}, \ldots, t_{qk})$. All individuals in the sample participate for some non-zero amount of time in maintenance activity, and hence this alternative constitutes the “outside good” in the MDCEV component of the model system.

Using the above notational preliminaries, we next discuss the structure of the residential location choice model component (Section 2.1.1), the time use model component (Section 2.1.2), and then highlight the joint nature of the two components (Section 2.1.3).

2.1.1 The Residential Location Component

The residential location component takes the familiar discrete choice formulation, as presented below:

$$ u_{qi}^* = \phi_q^i z_{qi} + \xi_{qi}, \text{ spatial unit } i \text{ chosen if } u_{qi}^* > \max_{d=1,2,\ldots,I} u_{qd}^* . \tag{1} $$

In the equation above, $u_{qi}^*$ is the indirect (latent) utility that the $q^{th}$ individual (as part of her/his household) obtains from locating in spatial unit $i$, $z_{qi}$ is a vector of ATE attributes corresponding to individual $q$ and spatial unit $i$ (such as land-use mix and measures of activity accessibility), and $\phi_q^i$ is a coefficient vector capturing individual $q$’s sensitivity to attributes in $z_{qi}$. We...
parameterize each element $l$ of $\varphi_q$ as $\varphi_{ql} = (\varphi_l + \lambda'_l w_{ql} + v_{ql} + \sum_k \omega_{qkl})$, where $w_{ql}$ is a vector of observed individual-specific factors (such as income and/or household size of individual $q$’s household) affecting sensitivity to the $l^{th}$ attribute in vector $z_{qi}$, and $v_{ql}$ and $(\sum_k \omega_{qkl})$ are individual-specific unobserved factors impacting individual $q$’s and her/his household’s sensitivity to the $l^{th}$ attribute in vector $z_{qi}$. $v_{ql}$ includes only those individual-specific unobserved factors that influence sensitivity to residential choice, while each $\omega_{qkl}$ ($k = 2, 3, \ldots, K$) includes only those individual-specific unobserved factors that influence both residential choice and time use in discretionary activity purpose $k$. For instance, consider an individual’s sensitivity to bicycling facilities around her/his household. The individual may have a higher sensitivity (than her or his observationally equivalent peer group) to bicycling facility density because of general auto disinclination. This auto disinclination may not, however, impact time use in activities. This would be captured in $v_{ql}$. Now, another unobserved individual factor may be fitness consciousness. This is likely to impact the sensitivity to bicycling facility density in residential choice (because better bicycling facilities are more conducive to bicycling activity) and also influence time invested in physically active leisure. This would be included in $\omega_{qkl}$ (more on this later). Finally, in Equation (1), $\varepsilon_{qil}$ is an idiosyncratic error term assumed to be identically and independently extreme-value distributed across spatial alternatives and individuals.

2.1.2. The Time Use Model Component
Designate the first alternative as maintenance activity, which is also the outside good that is always consumed. The rest of the ($K-1$) alternatives correspond to discretionary activities. Consider the following additive utility function form$^4$:

$$U_{qi}(t_q) = \frac{1}{\alpha_l} \psi_{qil} t_{q1}^\alpha_l + \sum_{k=2}^{K} \gamma_k \psi_{qkl} \ln \left( \frac{t_{qk}}{\gamma_k} + 1 \right).$$

(2)

In the above utility function, $t_q$ is the vector of time investments ($t_{q1}$, $t_{q2}$, $t_{qK}$) of individual $q$, or equivalently, the time spent in each activity purpose.$^5$ $U_{qi}(t_q)$ refers to the utility accrued to the individual due to time investment $t_q$ if s/he resides in spatial unit $i$. The term $\psi_{qkl}$

$^4$ Some other utility function forms were also considered, but the one below provided the best data fit. For conciseness, we do not discuss these alternative forms. The reader is referred to Bhat (2007) for a detailed discussion of alternative utility forms. The reader will also note the implicit assumption in the formulation below that there is utility gained from investing time in discretionary activities. This is a reasonable assumption since individuals have the choice not to participate in such activities. Also the reader will note that the inclusion of the maintenance activity type (the first alternative) allows the analyst to endogenously estimate the total amount of time invested in discretionary pursuits.

$^5$ The individual has the vector $t_q$ as the decision vector. The second through $K^{th}$ elements of $t_q$ can either be zero or some positive value (the first element of $t_q$ should be positive). Whether or not a specific $t_{kl}$ value ($k = 2, 3, \ldots, K$) is zero constitutes the discrete choice component, while the magnitude of each non-zero $t_{kl}$ value constitutes the continuous choice component. In the rest of this paper, we will use the terms “time investments” and “time use” interchangeably to refer to these discrete-continuous $t_{kl}$ values.
corresponds to the marginal random utility of one unit of time investment in alternative $k$ at the point of zero time investment for the alternative for the individual residing in spatial unit $i$ (as can be observed by computing $\partial U_{qi}(t_{q})/\partial t_{qk} \mid t_{qk}=0$). Thus $\psi_{qki}$ controls the discrete choice participation decision in alternative $k$ for individual $q$ residing in spatial unit $i$. We will refer to this term as the baseline preference for alternative $k$. The term $\gamma_k$ is a translation parameter that serves to allow corner solutions (zero consumption) for the “inside” alternatives $k = 2, 3, \ldots, K$ ($\gamma_k >0$). However, it also serves as a satiation parameter for these inside alternatives, with values of $\gamma_k$ closer to zero implying higher satiation (or lower time investment) for a given level of baseline preference (see Bhat, 2007). There is no $\gamma_1$ term for the first alternative in Equation 2 because it is always consumed. However, satiation effects in the consumption of this first alternative are captured through the exponential $\alpha$ parameter, which is bounded from above by a value of 1 (so that marginal utility decreases with increasing time investment in maintenance activity). The constraints that $\alpha \leq 1$ and $\gamma_k >0$ ($k = 2, 3, \ldots, K$) are maintained through appropriate parameterizations (see Bhat, 2007 for details)6.

To complete the model specification, we express the baseline parameters as functions of observed and unobserved attributes corresponding to the individual $q$ and spatial unit $i$ as follows: $\psi_{qki} = \exp(\beta_{q_k} z_{qi} + \theta_k x_{q} + e_{qki})$, where $z_{qi}$ is an ATE vector of observed attributes with individual-specific parameter vector $\beta_{q_k}$, and $x_{q}$ is a vector of pure individual sociodemographics with coefficient vector $\theta_k$.7 We further parameterize $\beta_{q_k}$ as follows: $\beta_{q_ki} = \beta_k + \Delta_{ki} s_{qi} + \eta_{qki}$, where $s_{qi}$ is a vector of observed individual-specific factors influencing the sensitivity of individual $q$ to the $l$th ATE attribute in vector $z_{qi}$ through the $\Delta_{kl}$ coefficient, and $\eta_{qkl}$ is a term capturing the impact of individual-specific and purpose-specific unobserved terms on the sensitivity to the $l$th ATE attribute in $z_{qi}$. $\eta_{qkl}$ is next partitioned into two components: $\sum_l (\pm \omega_{ql}) z_{qli} + \zeta_{qki}$. The $(\pm \omega_{ql}) z_{qli}$ terms are the common error components in the residential location and time investment model components, while $\zeta_{qki}$ is an idiosyncratic term assumed to be identically and independently standard type I extreme-value distributed across individuals, activity purposes, and spatial units.8

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6 The $\alpha$ and $\gamma$ parameters are subscripted only by activity purpose $k$ (unlike the $\psi$ parameters) because specification tests in our empirical analysis did not show statistically significant variation in these parameters based on individual and spatial unit-related characteristics.

7 Note that we are introducing the full vector $z_{qi}$ of ATE attributes in both the residential choice and time investment model components. In general, some ATE attributes will not impact residential choice (the corresponding element of $\varphi_q$ is zero for all $q$) and some will not influence time investments in activity type $k$ (the corresponding element of $\beta_{q_k}$ is zero for all $q$). Additionally, it is possible that ATE attributes have a mean effect of zero across individuals for residential choice and/or one or more activity purpose time investments, but have a significant distribution around the zero mean.

8 In our empirical analysis, we further partitioned $\zeta_{qki}$ into error components that generate covariance across discretionary activity purposes (for example, individuals who are generically inclined toward social activities may have a higher baseline preference than their observationally equivalent peers for both in-home and out-of-home
2.1.3. The Joint Model System

The specifications in the previous two sections may be collected and brought together in the following equation system:

\[
\begin{align*}
    u_{qi}^* &= \sum_i \left( \varphi_i + \lambda_i w_{ql} + v_{ql} \right) z_{qil} + \sum_k \left( \sum_l \omega_{qlk} \right) z_{qkl} + \xi_{qi} \\
    U_{qi}(t_q) &= \frac{1}{\alpha_i} \exp(\xi_{qil}) t_{qk} + \\
    &+ \sum_{k=2}^{k} \left\{ \gamma_k \exp \left[ \left( \sum_i \left( \beta_{hi} + \Delta_{hi} s_{qil} + \eta_{qkl} \right) z_{qil} \right) + \theta_{k} t_{q} + \left( \sum_l \left( \pm \omega_{qlk} \right) z_{qkl} \right) + \zeta_{qki} \right] \ln \left[ \frac{t_{qk}}{\gamma_k} + 1 \right] \right\}
\end{align*}
\] (3)

In the above equation system, the vectors \( z_{qi} \) and \( x_q \) do not appear for the first alternative in \( U_{qi}(t_q) \) because only utility differences matter. From the analyst’s perspective, individuals are choosing a residential location \( i \) and their time use profile by jointly maximizing \( u_{qi}^* \) and \( U_{qi}(t_q) \) subject to the budget constraint that \( \sum_k t_{q} = T_q \), where \( T_q \), as mentioned earlier, is the total time available to individual \( q \) for participation in maintenance and discretionary activities.

The joint nature of the model system arises because of the presence of the common unobserved \( \omega_{qkl} z_{qil} \) terms in the residential choice and time investment model components. More generally, the model system allows self-selection of individuals (based on their time-use preferences) into neighborhoods due to both observed and unobserved factors. In the context of observed factors, consider the situation where individuals who own bicycles are attracted toward residential neighborhoods with very good bicycling facilities. This can be reflected by including bicycle ownership as a variable in the \( w_{ql} \) vector that corresponds to the “bicycling facility density” variable in the \( z_{qi} \) vector of the residential choice equation. Individuals who own bicycles may also be more likely to spend time in physically active recreational pursuits, which can be accommodated by including bicycle ownership as a variable in the \( x_q \) vector in the time investment model component. If bicycle ownership were not included in the \( x_q \) vector for the physically active recreation purpose in the time investment equation, and if “bicycling facility density” were included as a variable in the \( z_{qi} \) vector of the same equation, the result could be an inflated estimate of the positive influence of bicycling facility density on time invested in physically active recreation. Similarly, consider an unobserved individual factor, such as fitness consciousness, that makes some individuals locate in areas with good bicycling facility and also participate more than their observationally equivalent peers in physically active recreation. Such factors are captured in the common \( \omega_{qkl} z_{qil} \) terms in the two equations. The ‘±’ sign in front of social activities). However, in our final specification, we did not find any such statistically significant covariance terms. Thus, for simplicity in presentation, we are imposing the restriction of zero covariances across discretionary activity purposes in the model structure presentation.
the $\omega_{qkl}z_{qil}$ terms in the time-investment model component indicates that the correlation in the unobserved factor $l$ may be positive or negative. If the sign is positive (negative), it implies that individuals who intrinsically prefer the ATE characteristics represented by $z_{qil}$ also have a higher (lower) generic preference for activity purpose $k$.

2.2 Model Estimation
The parameters to be estimated in the equation system (3) include the $\varphi_l, \lambda'_l, \Delta_{kl}$ and $\theta_k$ vectors, the $\gamma_k$ ($k = 2, 3, \ldots, K$), $\alpha_l$, and $\beta_{kl}$ scalars ($k = 2, 3, \ldots, K$; $l = 1, 2, \ldots, L$), and the variances of the stochastic components $v_{qil}$, $\eta_{qkl}$, and $\omega_{qkl}$ (all assumed to be normally distributed with variances $\sigma^2_{v}, \sigma^2_{\eta}, \sigma^2_{\omega}$, respectively) that represent random heterogeneity.

Let $\Omega$ represent a vector that includes all the parameters to be estimated, and let $\Omega_{-\sigma}$ represent a vector of all parameters except the variance terms. Also, let $g_q$ be a vector that stacks the $v_{qil}$, $\eta_{qkl}$, and $\omega_{qkl}$ terms, and let $\Sigma$ be a corresponding vector of standard errors. Define $a_{qi} = 1$ if individual $q$ resides in spatial unit $i$ and 0 otherwise. Then the likelihood function for a given value of $\Omega_{-\sigma}$ and $g_q$ may be written for an individual $q$ as:

$$L_q(\Omega_{-\sigma}) \big| g_q = \prod_{i=1}^{M_q} \left\{ \exp \left[ \sum_l \left( \varphi_l + \lambda'_l w_{qil} + v_{qil} + \sum_k \omega_{qkl} z_{qil} \right) \right] \right\} \times \left[ \prod_{h=1}^{M_q} e^{V_{qhi}} \right] \left( \sum_{h=1}^{K} e^{V_{qhi}} \right)^{M_q} (M_q - 1)! ,$$

where $M_q$ is the number of activity purposes in which the individual $q$ participates, and $|J_q|$ and $V_{qkl}$ are defined as follows:

$$|J_q| = \left( \prod_{h=1}^{M_q} c_{qh} \right) \left( \sum_{h=1}^{M_q} \frac{1}{c_{qh}} \right), \quad c_{qh} = \frac{1 - \alpha_q}{\gamma_q}, \quad c_{qh} = \frac{1}{\gamma_q + \gamma_h}, \quad (h = 2, \ldots, M_q)$$

$$V_{qil} = (\alpha_i - 1) \ln t^*_{qil}, \quad (i = 1, 2, \ldots, I),$$

$$V_{qkl} = \left( \sum_l \left( \beta_{kl} + \Delta_{kl} s_{qil} + \eta_{qkl} \right) z_{qil} + \theta_k x_q + \sum_l \left( \pm \omega_{qkl} \right) z_{qil} \right) - \ln \left( \frac{t^*_{qkl}}{\gamma_k} + 1 \right), \quad (k = 2, 3, \ldots, K).$$
In the above expression, the vector $\mathbf{t}^*_q = \{t^*_{q1}, t^*_{q2}, \ldots, t^*_{qk}\}$ represents the optimal (i.e., observed) time investments of individual $q$ in the various activity purpose.

Finally, the unconditional likelihood function can be computed for individual $q$ as:

$$L_q(\Omega) = \int_{\mathbb{R}^k} \left( L_q(\Omega|g_q) \right) dF(g_q|\Sigma),$$

where $F$ is the multidimensional cumulative normal distribution. The log-likelihood function can be written as: $L(\Omega) = \sum_q \ln L_q(\Omega)$. Simulation techniques are applied to approximate the multidimensional integral in Equation (4), and the resulting simulated log-likelihood function is maximized. Specifically, the scrambled Halton sequence (see Bhat, 2003) is used to draw realizations from the population normal distribution. In the current paper, we tested the sensitivity of parameter estimates with different numbers of scrambled Halton draws per observation, and found the results to be stable with as few as 100 draws. In the analysis, we used 125 draws per observation in the estimation.

### 3. THE DATA

#### 3.1 Data Sources

The primary source of data used for this analysis is the 2000 San Francisco Bay Area Travel Survey (BATS) designed and administered by MORPACE International Inc. for the Bay Area Metropolitan Transportation Commission (MTC). The survey collected information on all activity episodes (in-home and out-of-home) undertaken by individuals from over 15,000 households in the Bay Area for a two-day period (see MORPACE International Inc., 2002, for details on survey, sampling and administration procedures). Information characterizing the context (activity type, start and end times of the activity, and location of participation) of each activity episode was collected. Furthermore, data on individual and household socio-demographics was also obtained.

In addition to the 2000 BATS data, several other secondary data sources were used to derive spatial variables characterizing the activity-travel environment in the region. These include: (1) Land-use/demographic coverage data, obtained from the MTC, (2) The zone-to-zone motorized travel level of service files, obtained from MTC, (3) The Census 2000 population and Housing data summary files (SF1), (4) GIS layers of businesses (automotive businesses, shopping and grocery stores, medical facilities and personal services, food stores, sports and fitness centers, parks and gardens, restaurants, recreational businesses, and schools), obtained from the InfoUSA business directory, (5) GIS layers of bicycling facilities, also obtained from MTC, and (6) GIS layers of highway (interstate, national, state and county highways) network and local roadways (local, neighborhood, and rural roads) network, extracted from the Census 2000 Tiger files.

#### 3.2 Sample Formation

##### 3.2.1 Sample Extraction

Several steps were undertaken to generate the data for the current analysis. First, only adults (individuals 16 years or older) were selected to focus the analysis on the demographic segment that clearly exercises a choice in activity participation and time allocation. Second, only weekday
data was selected from the sample, since the focus of our analysis is on the time-use patterns on weekdays. Third, of the two-day activity diary data available for each individual, only one randomly chosen day was picked. In addition, we restricted the analysis to the time-use patterns of a single randomly chosen individual from each household. These decisions were made to keep the sample size manageable in the estimations, and also to avoid the problem of repeated data measurement from the same individual and/or household. Fourth, we selected only the non-work episodes for each individual. Finally, we confined the analysis to Alameda County in the San Francisco Bay area.

3.2.2 Activity Type Classification
The survey instrument used a 17-category activity purpose classification scheme, and collected information on the location of each activity episode. The activity purpose and location information was used to create a 13-category activity purpose typology in the current study to classify non-work episodes. The activity purposes are: (1) Maintenance (household chores, personal care, meal preparation, grocery shopping, and medical appointments), (2) In-home (IH) internet browsing, (3) Out-of-home (OH) volunteering (including religious and civic activity participation), (4) OH non-maintenance shopping (i.e., shopping for non-grocery purposes), (5) IH socializing, (6) OH socializing, (7) IH relaxing (resting, reading, listening to music, etc.), (8) IH recreation (hobbies, television viewing, etc.), (9) OH meals, (10) OH physically active pure recreation (episodes undertaken using non-motorized modes, and without any specific destination, such as walking, jogging, or bicycling around the neighborhood, are classified in this category), (11) OH physically inactive pure recreation (episodes undertaken using motorized modes, and without any specific destination, such as joy-riding around the block, belong to this category), (12) OH physically active recreation (exercising at the gym, playing tennis, etc.), and (13) OH physically passive recreation (going to the movies, opera show, etc.). The dependent variables in the time investment component of the model system correspond to the time allocated during the 24-hour survey weekday to each of the 13 activity purposes just identified. These variables are obtained by appropriate time aggregation across all episodes within each purpose category for each individual. The total time across all these 13 categories is considered exogenous, and the model focuses on residential choice and daily time-use in each of the 13 activity purpose categories.

3.2.3 Activity Travel Environment (ATE) Measures
The data from the secondary sources identified in Section 3.1 were used to compute a host of built environment measures for each traffic analysis zone (TAZ), including:
1. Zonal size and density measures, such as total population, population density, household density, density of employment by each of several employment categories, and dummy variables for central business district (CBD), urban, suburban, and rural areas (computed based on employment density). These attributes were obtained from the zonal land-use data file.
2. Zonal land-use structure variables, such as housing type measures (fractions of single family, and multiple family dwelling units), fractions of zonal area in residential and commercial land-uses, and land-use mix (see Bhat and Gossen, 2004 for a description of the land-use mix variable; the variable takes a value of zero for zones with only one type of land-use and a value of 1 for zones with equal distributions in area among residential, commercial, and other
land-uses). The zonal land-use structure variables were constructed from the zonal land-use data file.

3. Regional accessibility measures, such as shopping accessibility, recreational accessibility, and employment accessibility. These are Hansen-type (see Fotheringham, 1983) accessibility measures computed from the zonal land-use and level-of-service data.

4. Zonal demographics, such as average household size and median household income. These demographic measures were derived from the Census 2000 population and housing data summary file (SF1).

5. Zonal ethnic composition measures, constructed as fractions of Caucasian, African-American, Hispanic, Asian, and other ethnicity populations, also derived from the Census 2000 population and housing data summary file.

6. Zonal activity opportunity variables, such as activity center intensity (i.e., the number of business establishments per square mile) and density (i.e., the number of business establishments per square mile) for each of the following activity types (extracted from the InfoUSA business establishments data): (a) maintenance (grocery stores, gas stations, food stores, car wash, automotive businesses, banks, medical facilities, etc.), (b) physically active recreation (fitness centers, sports centers, dance and yoga studios, parks, gardens, etc.), (c) Physically inactive recreation (theatres, amusement centers, arcades, etc.), and (d) eat-out (restaurants and eateries).

7. Zonal transportation network measures, such as highway density (miles of highway facilities per square mile), bikeway density (miles of bikeway facilities per square mile), and local roadway density (miles of roadway density per square mile). These variables were extracted from the GIS layers of bikeways and roadways.

All of these ATE attributes were merged with the activity time-use and individual/household demographic data to form a comprehensive database suitable for modeling residential choice and time-investment decisions.

3.3 Sample Description
The final estimation sample consists of 2793 individuals residing in Alameda County in the San Francisco Bay area. Each individual has the choice of residing in any one of 236 zones in Alameda County, which is modeled using a standard discrete choice formulation. In this residential choice model, we do not undertake a sampling scheme to reduce the number of alternatives because our overall residential choice model does not correspond to a simple multinomial logit model (rather there is mixing in the model).

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9 We consciously chose not to consider some other ATE variables that have typically been considered in residential choice models. These include average (or median) housing costs in a zone, and commute-related variables (travel times and costs from a candidate residential zone to the work place zone of the individual). The primary reason for not including these variables in our model is that earlier studies have underscored the joint nature of residential and workplace decisions (see Siegel, 1975, Simpson, 1980, Waddell, 1993a, and Merriman, 1994), and highlighted the endogeneity of housing costs in residential choice decisions (see Guevara and Ben-Akiva, 2006, and de Palma et al., 2005). In our study, which emphasizes the need to consider residential choice as endogenous to activity time-use decisions, we did not want to be inconsistent in spirit and summarily ignore the potential endogeneity of work place choice and housing costs in residential choice. At the same time, we wanted to restrict our attention here to just two decision processes, given that this is the first research attempt to consider a joint MNL-MDCEV structure. Consequently, we do not include commute-related characteristics and housing costs as exogenous variables in the model. We leave the extension of our model structure to explicitly incorporate the endogeneity of commute characteristics and housing costs to future research.
The time-use of individuals in each of 13 activity purposes constitutes the dependent variables of the MDCEV component of the model system. Each (and all) of the 2793 individuals participated in maintenance activity. The frequency distribution of individuals based on the number of discretionary activity types they participated in is provided in Table 1. As can be observed, 32.6% of individuals did not participate in any discretionary activity, while the remaining 67.4% of individuals participated in at least one type of discretionary activity. Of the individuals who participated in at least one type of discretionary activity, 51.2% of individuals participated in multiple types of discretionary activities. Overall, the results clearly illustrate the high prevalence of participating in multiple non-work activity purposes on a single weekday, providing strong support for the use of the MDCEV model for time-use analysis.

4. EMPIRICAL ANALYSIS

4.1 Variable Specification
Several types of variables were considered in the joint modeling system. These included: (1) household socio-demographics (household size, household composition and family structure, vehicle and bicycle ownership, ethnicity, income, whether or not household owned or rented its residence, dwelling type, availability of internet, etc.), (2) individual demographics and employment characteristics (age, license holding to drive, employment status, number of hours of work on the day, physical disability status etc.), (3) contextual variables such as season of the year, day of the week, rain-fall (and the amount of rain fell) in the day, temperature (minimum and maximum values in the day, and the variation in the day), (4) a host of activity-travel environment variables (discussed in section 3.2.3), and (5) the interactions of the activity-travel environment variables with household and individual socio-demographics.

The final variable specification was obtained based on a systematic process of eliminating variables found to be statistically insignificant, and parsimony in representation. The specification was additionally guided by intuitive considerations, and results from earlier studies. In the next section (Section 4.2), we discuss the results of the residential location choice model component. Section 4.3 presents the results of the activity time-use model component. Section 4.4 discusses self-selection effects, and Section 4.5 focuses on likelihood-based measures of data fit.

4.2 Residential Location Choice Model Component Results
Table 2 presents the estimation results for the residential location choice component of the joint model. The first set of variables in Table 2 corresponds to zonal size and density measures. The coefficient on the logarithm of number of households in the zone has the expected positive sign, indicating that individuals are more likely to locate in zones with a large number of housing opportunities. The parameter on this variable is between 0 and 1, as should theoretically be the case on the size measure (see Bhat et al., 1998 and Daly, 1982). The household density effects indicate that individuals whose households do not have seniors (of age > 65 years) and do not have young children (≤ 15 years of age) are likely to locate in zones with high household density, while those with seniors and young children prefer low density neighborhoods. The latter result may be a reflection of the “urban-to-suburban flight” trend of households that have young children, as has been documented in several sociological studies (Birch, 2005, Lee and Guest, 1983, Marans, 1979, and Waddell and Nourzad, 2002). This flight trend has been attributed to the safety and neighborhood quality related preferences of households having children. The employment density coefficients indicate that low income individuals (annual
household income < $35,000) and African-American households tend to locate themselves in areas with high employment density, while those who have high incomes, who have children (15 years of age or younger) in their household, and who are Caucasians prefer areas of low employment density.

The zonal land-use structure variable effects indicate that individuals shy away from zones with a high fraction of commercial land area and with good land use-mix diversity. The effect of land-use mix variable is interesting, and suggests that individuals do, in general, prefer homogenous, exclusive enclaves, for their home location.

Our results did not indicate any statistically significant effects of the regional accessibility measures. The zonal demographic variables show strong effects on residential choice. In particular, there is a clear indication of residential clustering of individuals based on income levels, household size, and age. Such clustering trends have long been documented in the residential analysis literature (see Waddell, 1993b, and Waddell, 2006).

The zonal race variables, when interacted with individual race, show clear evidence of racial clustering, with the effect being strongest for African-American individuals and weakest for Caucasian individuals. The clustering effect for Asian and Hispanic individuals is somewhere between those for African-Americans and Caucasians, with Asians having a higher tendency to cluster than Hispanics. The racial clustering may be a result of social networking, race-specific preferences for unobserved zonal characteristics, discrimination in the housing market place, or some combination of these. Disentangling the possible reasons for racial clustering is an important avenue for further research.

The next set of variables in Table 2 represents the opportunities offered by a zone for different types of activities. The positive coefficient on the number of schools in a zone indicates that households are more likely to locate in zones with better schooling opportunities. Interestingly, we did not find any statistically significant variation in the effect of number of schools based on the presence (or number) of children in the household. This is, in part, because a high fraction of individuals in our sample have children in their households, leading to inadequate variation in the children-related variables. The next variable, “number of physically active recreation centers such as fitness centers/gymnasiums, sports centers, dance and yoga centers”, suggests a positive impact of the zonal level opportunities for physically active recreational activity pursuits on a household’s preference to reside in that zone. The final variable under zonal activity opportunities indicates the preference of individuals with bicycles in their households to locate in zones with a higher number of natural recreation opportunities, perhaps due to the generally outdoor activity inclination of individuals with bicycles. Admittedly, none of these zonal activity opportunity variables have a highly statistically significant impact on residential choice, but they are included so that their effects can be explored more carefully in future residential choice studies that use a neighborhood-level scale for measuring these opportunity attributes (rather than the aggregate zone-level scale in the current study).

The effects of the zonal transportation network measures show that individuals with seniors (age > 65 years) in their household are likely to stay away from zones with high highway density. This is possibly a reflection of the preference for relatively quiet, retirement-like, communities amongst households with seniors (though this effect is not very statistically significant). The influence of street block density lends reinforcement to the stereotype of high income households choosing to locate in suburban-like, low density, sprawling, communities away from the “inner city” neighborhoods. Finally, in the set of zonal network measures, the
results show that, without exception, households prefer zones with high bicycling facility density, though this effect is particularly strong among individuals who own several bicycles.

4.3 Time-use Analysis Component Results
The final specification results of the activity time-use component of the joint model (i.e., the MDCEV modeling component) are presented in Table 3. Each row corresponds to a parameter or a variable, and each column corresponds to a specific activity purpose category. The maintenance activity purpose serves as the base category for all variables. In addition, a blank (i.e., no parameter estimate) for a variable for an alternative implies that the alternative also constitutes the base category for that variable.

4.3.1 Baseline Preference Constants
The baseline preference constants (see first row of Table 3) do not have any substantive interpretations. They capture generic tendencies to participate in each discretionary activity type category as well as accommodate the range of continuous independent variables in the model. However, all the baseline preference constants are negative, which is reflection of the fact that all individuals participate in maintenance activity (the “outside” good), while this is not the case for the remaining discretionary activity purposes.

4.3.2 Satiation/Translation Parameters
The satiation parameter estimate of $\alpha_1$ for the maintenance activity (i.e., the outside good) alternative is 0.253, with a t-statistic value of 24.82 for the null hypothesis that $\alpha_1 = 1$. This indicates the presence of strong satiation effects in the time investment on maintenance activities. That is, the marginal utility of time investment in maintenance activity drops rapidly with increasing maintenance time.

The translation parameters ($\gamma_k$) for the discretionary activity purposes are presented in the second row of Table 3. As discussed in Section 2.1.2, the $\gamma_k$ parameters allow corner solutions (i.e., zero time investment in the discretionary activity purposes) as well as serve as satiation parameters. The magnitude of the $\gamma_k$ parameter for any activity purpose $k$ is inversely associated with satiation for that activity purpose. Thus, a value closer to zero for activity purpose $k$, in general, implies higher satiation for activity purpose $k$ (and, therefore, lower time investment in activity purpose $k$), while a high value for an activity purpose implies lower satiation for that activity purpose (or, equivalently, higher time investment in that activity purpose). The results show high satiation effects (low durations) for time investments in out-of-home (OH) shopping, OH meals, OH physically activity pure recreation, and OH physically inactive pure recreation. On the other hand, the results indicate the low satiation effects (high durations) for time investments in in-home (IH) relaxation and IH recreation. The satiation effects for the remaining activities are between these two extremes. These results are consistent with intuitive expectations.

4.3.3 Household Sociodemographics
Among the household sociodemographic variables, the effect of household structure on the baseline utilities indicates that individuals living alone are more likely to partake in IH socializing, and less likely to participate in IH relaxing, compared to individuals not living alone. This is perhaps a reflection of the higher need to socialize and interact with other individuals.
when living alone (see Yamamoto and Kitamura, 1999 for similar results). The specific preference for in-home social activities over out-of-home social activities among individuals living alone is interesting, and needs further exploration in future studies.

The race variable effects show that Caucasians have a higher baseline preference for OH meal activity relative to other races. Other than this, there are no statistically significant differences among races in time use patterns. This is in contrast to studies of weekend activity time-use that show that Caucasians are not only more likely than non-Caucasians to participate in OH meal activity, but also in all other kinds of OH recreational activities (see Kapur and Bhat, 2007; and Bhat et al., 2006). This points to more heterogeneity in time-use among races during the weekends relative to weekdays. The next household sociodemographic attribute is the availability of internet at home, which, as expected, is associated with a higher tendency of individuals to invest time on internet use.

Household income also impacts time investment patterns in activity purposes. Specifically, individuals in low income households have a higher baseline preference for in-home relaxing and in-home recreation relative to individuals in middle and high income households, while those in middle and high income households are more likely to participate in out-of-home meals and out-of-home recreation pursuits. These results are consistent with the higher consumption potential of services and out-of-home recreation facilities of individuals in higher income earning households.

Finally, within the set of household sociodemographic variables, the coefficient on the number of motorized vehicles suggests a positive association between high motorized vehicle ownership and an inactive life style (i.e., lower preference for physically active pure recreational travel as well as physically active out-of-home recreational activities). On the other hand, high bicycle ownership is associated with a lower preference for in-home relaxing and a higher preference for physically active pure recreational travel. The latter result is quite reasonable. Households who own more bicycles are likely to be more outdoor-oriented by nature, as discussed in the residential location choice model results in Section 4.2. In that section, we also pointed to the increased likelihood of individuals with high bicycle ownership to locate themselves in zones with good bicycling facilities. Thus, by including the bicycle ownership variable in the time-use model component, we are capturing residential self-selection effects due to bicycle ownership. Specifically, if bicycle ownership were not included in the time-use model component, it would lead to an over-inflated effect of the potential to encourage physically active pure recreation participation by designing neighborhoods with good bicycling facility (as we found to be the case in our empirical analysis when we removed the bicycle ownership variable from the time-use model specification). Overall, the results pertaining to vehicle ownership emphasize the transportation-public health connection. That is, our analysis suggests that policies and educational campaigns aimed at reducing motorized vehicle ownership and increasing bicycle ownership not only can lead to traffic congestion alleviation, but can also play an important role in improving public health.

An interesting point is in order here regarding the effects of household sociodemographics. In earlier studies of weekend time-use, the presence and number of children

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10 See Pinjari et al. (2007) for a similar finding on residential self selection effects due to bicycle ownership in a commute mode choice model. Other studies that accommodate such self selection effects due to observed factors include, for example, Bagley and Mokhtarian (2002), Bhat and Guo (2007), Cao et al. (2006), Guo et al. (2007), Handy et al. (2005), Kitamura et al. (1997), Khattak and Rodriguez (2005), and Schwanen and Mokhtarian (2004, 2005a).
has been found to play an important role in adult time-use decisions. For instance, Bhat et al. (2006) indicate that individuals in households with children have a high baseline preference for out-of-home recreation and pure recreational pursuits, and a lower preference for in-home leisure activities. They attribute these effects to a stronger need to have a change from caring for children in-home and the propensity to participate with young children in outdoor pursuits. In contrast, we did not find any statistically significant effect of children on time-use patterns on weekdays. This is perhaps because of weekday work- and school-related activities, because of which there is less of a need to have a change from caring for children in-home and less opportunity for joint participation between adults and children.

4.3.4 Individual Sociodemographics and Employment Characteristics

Among the individual sociodemographics, the coefficient on the male dummy variable highlights the role of gender in weekday discretionary time-use. Women are more inclined to participate in volunteering activities, while men participate more in internet use, in-home relaxing, in-home recreation, out-of-home meals, and OH physically active recreation. The higher participation of men in relaxing/meal/recreational pursuits also implies, because of the fixed time constraint across all discretionary and maintenance activities, that men participate less in maintenance activity. That is, women have more responsibility for household maintenance activity, a result consistent with the findings of several earlier studies (see, for example, Chen and Mokhtarian, 2006, Gossen and Purvis, 2005, Goulias, 2002, Levinson, 1999, and Srinivasan, 2005).

Several different functional form specifications were attempted for the age-related variables, including a continuous variable, spline variables that allow piece-wise linear effects of age, and age dummy variables. After extensive testing, the best results were obtained using dummy variables for age less than 30 years, age between 30-65 years, and age greater than 65 years. The age variables are introduced in Table 3 with the age category between 30 to 65 years as the base. The results show that young adults (16-29 years) are more likely than other adults to participate in internet browsing and in all discretionary activities other than OH volunteering, OH shopping and physically active recreational pursuits (see Yamamoto and Kitamura, 1999 and Bhat et al., 2006 for similar results). On the other hand, older adults (>65 years) have a higher baseline preference for out-of-home volunteering activities.

The individual mobility-related variables indicate that, in general, individuals with a driver’s license have a higher preference (than those without a driver’s license) to participate in out-of-home discretionary activities, while those who are physically challenged are less inclined to participate in out-of-home discretionary activities. These are clearly manifestations of enhanced mobility to access activities (in the case of having a driver’s license) and mobility constraints (in the case of being physically challenged).

The effect of employment can be discerned from the coefficients on the employment dummy and work duration variables in Table 3. Overall, employed individuals have a lower baseline preference (relative to unemployed individuals) for all in-home activities other than relaxing. They also have a higher preference for IH relaxing, OH meal activities, and a lower preference for other OH activities. Further, among those employed, individuals who work longer have a particularly high preference for IH relaxing and a low preference for OH discretionary activities. These results reflect time constraints that make employed individuals (particularly those working long hours) spend more of their non-work time on maintenance activities (the base category), in-home relaxation, and out-of-home meals.
4.3.5 Day of Week and Seasonal Variables

The day of week effects reveal the higher inclination of individuals to participate on Fridays (relative to other weekdays) in such discretionary activities as OH shopping, socializing (in-home as well as out-of-home), IH relaxing, IH recreation, OH meals, and OH physically inactive recreation. The results also show the lower preference for pure recreation activities during the summer season, and for all kinds of recreation and IH relaxing activities in the Fall. These seasonal variations need further exploration in future studies. Our attempts to include weather-related factors (such as rainfall and temperature) were not successful.

4.3.6 Activity-Travel Environment Attributes

Among the activity-travel environment attributes, the parameters of the employment density variable indicate that individuals residing in high employment density neighborhoods are more likely to spend their leisure time at home than out-of-home. This may possibly be a reflection of traffic congestion and mobility problems in areas with high employment density that leads to higher in-home discretionary activity participation.

The variables representing the opportunities offered by a zone for various types of activities reveal the higher likelihood of participation in OH meal activity (OH physically inactive recreation) in areas with several eat-out centers (physically inactive recreational centers) per square mile. Also, individuals who own one or more bicycles, and live in areas with a high intensity of availability of sports/fitness centers, have a higher baseline preference for OH physically active recreational pursuits. This reflects the interaction effect of owning a bicycle and the presence of sports/fitness centers on OH physically active participation, perhaps because individuals who own a bicycle are also fitness-conscious.

The positive impact of the bicycling facility density variable on OH physically active pure recreation suggests that better bicycling facilities do lead to higher participation rates in physically active pursuits such as walking, biking, or jogging. The coefficient on the interaction of the bicycling facility density variable with household vehicle availability indicates that vehicle ownership moderates down the positive impact of bicycling facilities on OH physically active pure recreation. This finding further adds to the evidence on the association of higher vehicle ownership with physically inactive lifestyles (see Section 4.3.3).

4.4 Self-Selection Effect

In our model specifications, we explored unobserved heterogeneity in the effects of several variables in the residential and time-investment model components [corresponding to the \( v_{ql} \) and \( \eta_{qkl} \) terms in Equation (3)]. None of these effects turned out to be statistically significant. However, we did find a statistically significant standard deviation of the unobserved component associated with the bicycling facility density variable that was common to both the residential choice component and the OH physically active recreation baseline utility in the time-investment model component (corresponding to \( \omega_{qkl} \) in Equation 3). The estimated standard deviation was 0.042 with a t-statistic of 2.31. The sign corresponding to this variable in the \( \pm \omega_{qkl}z_{qil} \) term of the time-use component was positive. The implication is that there are unobserved individual factors (such as fitness consciousness that cannot be completely captured by such observed variables as bicycle ownership) that make individuals locate in areas with good bicycling facilities and also lead to a high preference for physically active recreation. That is, people who are predisposed to physically active lifestyles tend to self-select themselves into zones with very good bicycling
facility density for their residence. If this residential sorting is not accounted for, and the time investment model includes the bicycling facility density variable, the result could be a spurious finding that bicycling facility density causes higher participation levels in OH physically active recreation. In fact, in our model accommodating self-selection, the impact of bicycling facility density on the OH physically active recreation category was statistically insignificant (see the blank cell corresponding to the OH physically active recreation column for the bicycling facility density variable in Table 3). However, when we estimated an independent time-investment model without considering self-selection, the coefficient on bicycling facility density corresponding to the OH physically active recreation activity purpose was positive and highly significant. Clearly, this shows the danger of estimating the effects of the activity-travel environment on activity-travel choices without considering self-selection effects. More generally, it is important to model residential location choice and activity time-use choices in a joint framework to obtain the “true” effects of ATE attributes on activity time-use choices.

### 4.5 Overall Likelihood-based Measures of Fit

The log-likelihood value at convergence for the final joint multinomial logit (MNL)–multiple discrete-continuous extreme value (MDCEV) model is –42759.43. The corresponding value for the independent MNL and MDCEV models with no allowance for residential self selection due to unobserved decision-maker attributes is –42763.84. The likelihood ratio index for testing the presence of residential self selection due to unobserved factors is 8.82, which is larger than the critical chi-square value with 1 degree of freedom at a level of significance greater than 99.5% (the 1 degree of freedom corresponds to the standard deviation of the common error component related to the coefficient on bicycling facility density variable). Further, the log-likelihood value is –44127.17 for the independent MNL and MDCEV model system with (a) equal probability for each of the 236 spatial alternatives in the residential location MNL model and (b) only constants in the baseline preference terms, and the satiation and translation terms, in the activity time-use MDCEV model. The likelihood ratio index for testing the presence of exogenous variable effects and residential self selection effects is 2735.48, which is substantially larger than the critical chi-square value with 88 degrees of freedom at any reasonable level of significance. This clearly underscores the value of the model estimated in this paper to predict residential location and time-use choices of individuals as a function of relevant exogenous variables and accommodating residential self-selection effects.

### 5. DEMONSTRATION OF MODEL APPLICATION

The estimates of the activity time-use component of the joint model system can be applied to predict the changes in time-use patterns due to changes in sociodemographic characteristics and ATE attributes. In this section, we demonstrate the application of the model by studying the effects on aggregate time-use patterns of changes in two sociodemographic characteristics and two ATE attributes. The sociodemographic changes correspond to a decrease in household vehicle ownership for each individual by 1 (except for zero car households whose car ownership level is left unchanged), and an increase in household bicycle ownership for each individual by 1. The ATE attribute changes include a ten-fold increase in physically inactive recreation centers per square mile in each TAZ, and a ten-fold increase in bicycling facility density in each TAZ. To examine the effects of each of these changes, we computed the percentage change in the aggregate time investment in the five recreational activity categories (i.e., IH recreation, OH
physically active pure recreation, OH physically inactive pure recreation, OH physically active recreation, and OH physically inactive recreation) from before- to after- the change.\textsuperscript{11}

Table 4 presents the results. A “--” entry in the table indicates changes less than 1%. Four important observations may be made from this table. First, and as indicated earlier in Section 4.3.3, policies and educational campaigns aimed at reducing motorized vehicle ownership and increasing bicycle ownership not only can lead to traffic congestion alleviation, but can also play an important role in improving public health (as evidenced by the increase in overall time-use in the OH physically active recreation categories in the table). Second, the activity-travel environment can be used as a tool to influence individual activity time-use patterns. In the table, this is reflected in the change in OH physically inactive recreation time due to the change in the number of physically inactive recreation centers per square mile, and the change in OH physically active pure recreation time due to the change in bicycling facility density. Third, while the results show evidence that the activity-travel environment can be engineered to influence activity time-use, the results also show that the ability to do so is very limited. Specifically, the changes in activity time-use are highly inelastic to changes in ATE attributes. Fourth, while the percentage changes in activity time-use due to changes in sociodemographics and changes in ATE attributes are not strictly comparable (because the sociodemographic variables correspond to a change by 1 to ordinal variables, while the ATE attributes correspond to a 1000% increase to continuous variables), there is an indication that sociodemographics play a far more dominant role in determining activity time-use than do ATE attributes.

6. CONCLUSIONS
This study contributes to both the integrated land use-transportation modeling and activity-based analysis literature by presenting a joint model of residential location choice and activity time-use choices. This modeling system considers a comprehensive set of activity-travel environment (ATE) variables and sociodemographic variables as potential determinants of weekday time-use choices. The model formulation takes the form of a joint mixed Multinomial Logit–Multiple Discrete-Continuous Extreme Value (MNL–MDCEV) structure that (a) accommodates differential sensitivity to the ATE attributes due to both observed and unobserved individual and household attributes, and (b) controls for the self selection of individuals into neighborhoods based on observed and unobserved activity time-use preferences. To our knowledge, the analysis in this paper presents the first instance of the formulation and application of such a unified econometric methodological framework for jointly modeling (a) residential location choice (an unordered multinomial discrete choice variable), and (b) activity time-use choice (a multivariate discrete-continuous choice variable), while also accounting for observed and unobserved heterogeneity in the choice processes. The joint model system is estimated on a sample of 2793 households and individuals residing in Alameda County in the San Francisco Bay Area.

The model results offer several insights regarding individual time-use choices. First, both household/individual sociodemographics, and ATE attributes, are influential in shaping activity time-use choices. However, sociodemographic characteristics are more influential in shaping activity time-use behavior than are the ATE attributes. Second, there are significant observed factors contributing to residential self selection based on activity time-use preferences. For

\textsuperscript{11}We do not show the effects on other activity purposes in the table to reduce clutter and because these other effects are rather small. Also, to keep the constrained optimization-based prediction process manageable (see Bhat, 2005 for the prediction process details), we randomly chose 150 individuals from the estimation sample for the prediction analysis.
instance, individuals with bicycles locate themselves into neighborhoods with good bicycling facilities. These same individuals also have a preference for physically active pure recreation pursuits. Ignoring this effect of bicycle ownership would lead to an inflated estimate of the effect of bicycling facility density on time invested in physically active recreation. Similarly, high income households locate in neighborhoods with low employment density and low street block density. Individuals from such high income households also have a preference for out-of-home recreational activities/travel. Thus, ignoring income effects in activity time-use choices can lead to a spuriously estimated negative effect of employment density and street block density on out-of-home recreational activities/travel. Overall, even if there are no common unobserved factors influencing residential choice and activity time-use choices, the results suggest that it behooves the analyst to estimate an independent residential choice model so that any observed demographic factors impacting the sensitivity to ATE attributes in residential choice can be considered in the time-use model. In this way, one can reduce the possibility of “corrupt” inferences regarding the impact of ATE attributes on time-use choices. Of course, another reason to model both residential choice and time-use choice, even in the absence of common unobserved factors, is the fact that ATE attributes impact both these choices. Thus, policy decisions regarding changes in ATE characteristics have to be evaluated in the context of spatial relocations as well as time-use shifts to obtain a comprehensive picture of the changes due to ATE-related policies. Third, people who are predisposed to physically active lifestyles appear to self-select themselves into zones with very good bicycling facility density for their residence. Ignoring such unobserved factors resulted, in our current empirical study, in a spuriously estimated positive effect of the bicycling density variable on physically active recreational activity participation. This result shows the danger of estimating the effects of the activity-travel environment on activity-travel choices without considering self-selection effects due to unobserved factors. Fourth, ATE attributes such as employment density, activity opportunity density/intensity, and bicycling facility density have statistically significant impacts on activity time-use decisions even after controlling for residential sorting effects due to both observed and unobserved factors. Fifth, our results emphasize the transportation-public health connection. That is, our analysis suggests that policies and educational campaigns aimed at reducing motorized vehicle ownership and increasing bicycle ownership, when combined with better provision of bicycling facilities, not only can lead to traffic congestion alleviation, but can also play an important role in improving public health through increased investment of time in physically active pursuits.

To summarize, our results indicate that activity-travel environment attributes are not “completely” exogenous in activity time-use decisions. Households and individuals locate themselves in ATEs that are consistent with their lifestyle preferences, attitudes, and values. In other words, households and individuals make residential location and activity time-use decisions jointly as part of an overall lifestyle package. Nevertheless, the findings from this research suggest that modifying the activity-travel environment can bring about changes in activity time-use patterns, as demonstrated in the section on model application.

Of course, the research in the current paper can be extended and enhanced by (1) undertaking the analysis at a more disaggregate spatial level of analysis than traffic analysis zones (though such a move also promises to raise some very important computational challenges), (2) applying the methodology developed in this paper to richer data sets with attitudinal variables that may further enhance our understanding of the relationship of ATE
attributes on activity-travel dimensions, and (3) extending the analysis to include work location choice, and explicitly considering housing price endogeneity in residential choice analysis.

ACKNOWLEDGEMENTS
The authors would like to thank Lisa Macias for her help in typesetting and formatting this document. The second author would like to acknowledge the support of an International Visiting Research Fellowship and Faculty grant from the University of Sydney. This research was partially funded by Environmental Protection Agency Grant R831837.
REFERENCES


Handy, S., Cao, X., and Mokhtarian, P. (2005) Correlation or causality between the built environment and travel behavior? Evidence from Northern California. Transportation Research D, 10 (6), 427-444.


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Table 1. Distribution of Individuals by Number of Discretionary Activity Purposes Participated In

<table>
<thead>
<tr>
<th>Number of discretionary activity types participated in</th>
<th>Number of individuals participated</th>
<th>(% of individuals participated)</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>910</td>
<td>(32.6%)</td>
</tr>
<tr>
<td>1</td>
<td>919</td>
<td>(32.9%)</td>
</tr>
<tr>
<td>2</td>
<td>610</td>
<td>(21.8%)</td>
</tr>
<tr>
<td>3</td>
<td>241</td>
<td>(8.6%)</td>
</tr>
<tr>
<td>4</td>
<td>88</td>
<td>(3.2%)</td>
</tr>
<tr>
<td>5</td>
<td>20</td>
<td>(0.7%)</td>
</tr>
<tr>
<td>6</td>
<td>5</td>
<td>(0.2%)</td>
</tr>
<tr>
<td>Total</td>
<td>2,793</td>
<td>100%</td>
</tr>
</tbody>
</table>
Table 2. Estimation Results of the Residential Location Choice Component of the Joint Model

<table>
<thead>
<tr>
<th>Variables</th>
<th>Parameter</th>
<th>t-stat</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Zonal size and density measures (including demographic interactions)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Logarithm of number of households in zone</td>
<td>0.836</td>
<td>16.97</td>
</tr>
<tr>
<td>Household density (#households per acre x 10^{-1})</td>
<td>0.238</td>
<td>3.61</td>
</tr>
<tr>
<td>Interacted with presence of seniors in household</td>
<td>-0.736</td>
<td>-6.41</td>
</tr>
<tr>
<td>Interacted with presence of children (of age &lt; 5 years) in household</td>
<td>-0.392</td>
<td>-2.30</td>
</tr>
<tr>
<td>Interacted with presence of children (of age 5 to 15 years) in household</td>
<td>-0.216</td>
<td>-2.23</td>
</tr>
<tr>
<td>Employment density (#jobs per acre x 10^{-1})</td>
<td>-0.006</td>
<td>-1.02</td>
</tr>
<tr>
<td>Interacted with household income less than $35,000 per annum</td>
<td>0.196</td>
<td>2.38</td>
</tr>
<tr>
<td>Interacted with household belonging to the African American race</td>
<td>0.114</td>
<td>1.81</td>
</tr>
<tr>
<td>Interacted with presence of children (of age 15 years or younger) in household</td>
<td>-0.069</td>
<td>-1.81</td>
</tr>
<tr>
<td>Interacted with household belonging to the Caucasian race</td>
<td>-0.067</td>
<td>-1.48</td>
</tr>
<tr>
<td><strong>Zonal land-use structure variables (including demographic interactions)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fraction of commercial land area</td>
<td>-0.177</td>
<td>-1.04</td>
</tr>
<tr>
<td>Land-use mix</td>
<td>-0.216</td>
<td>-1.98</td>
</tr>
<tr>
<td><strong>Zonal demographics (including demographic interactions)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Absolute difference between zonal median income and household income ($ x 10^{-1})</td>
<td>-0.022</td>
<td>-15.29</td>
</tr>
<tr>
<td>Absolute difference between zonal average household size and household size</td>
<td>-0.457</td>
<td>-7.87</td>
</tr>
<tr>
<td>Fraction of senior (age &gt; 65) population interacted with presence of senior adult in Household</td>
<td>2.760</td>
<td>4.67</td>
</tr>
<tr>
<td><strong>Zonal race composition measures (including demographic interactions)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fraction of African-American population interacted with African-American dummy variable</td>
<td>3.931</td>
<td>9.35</td>
</tr>
<tr>
<td>Fraction of Asian population interacted with Asian dummy variable</td>
<td>3.241</td>
<td>8.06</td>
</tr>
<tr>
<td>Fraction of Caucasian population interacted with Caucasian dummy variable</td>
<td>1.912</td>
<td>13.48</td>
</tr>
<tr>
<td>Fraction of Hispanic population interacted with Hispanic dummy variable</td>
<td>2.130</td>
<td>3.09</td>
</tr>
<tr>
<td><strong>Zonal activity opportunity variables (including demographic interactions)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Number of schools in the zone</td>
<td>0.020</td>
<td>1.90</td>
</tr>
<tr>
<td>Number of physically active recreation centers such as fitness centers, sports centers, dance and yoga studios</td>
<td>0.080</td>
<td>1.56</td>
</tr>
<tr>
<td>Number of natural recreational centers such as parks, gardens, etc. interacted with Number of bicycles in the household</td>
<td>0.012</td>
<td>1.03</td>
</tr>
<tr>
<td><strong>Zonal transportation network measures (including demographic interactions)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Highway density (miles of highway per square mile) interacted with presence of seniors (age &gt; 65) in household</td>
<td>-0.040</td>
<td>-1.01</td>
</tr>
<tr>
<td>Street block density (number of blocks per square mile x 10^{-1}) interacted with household income greater than $90,000 per annum</td>
<td>-0.044</td>
<td>-4.12</td>
</tr>
<tr>
<td>Bicycling facility density (miles of bike lanes per square mile)</td>
<td>0.028</td>
<td>3.03</td>
</tr>
<tr>
<td>Interacted with number of bicycles in the household</td>
<td>0.016</td>
<td>3.93</td>
</tr>
</tbody>
</table>
### Table 3. Estimation Results of the Time-use Component of the Joint Model

<table>
<thead>
<tr>
<th></th>
<th>OH Volunteering</th>
<th>IH Internet use</th>
<th>OH Shopping</th>
<th>IH Socializing</th>
<th>OH Socializing</th>
<th>IH Relaxing</th>
<th>IH Recreation</th>
<th>OH Meals</th>
<th>OH Physically Active Pure Recreation</th>
<th>OH Physically Inactive Pure Recreation</th>
<th>OH Physically Active Recreation</th>
<th>OH Physically Inactive Recreation</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Baseline Preference Constants</strong></td>
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<tr>
<td></td>
<td>-6.751 (-33.93)</td>
<td>-9.321 (-16.73)</td>
<td>-5.985 (-29.52)</td>
<td>-7.863 (-29.26)</td>
<td>-6.668 (-39.49)</td>
<td>-5.669 (-28.80)</td>
<td>-5.667 (-29.09)</td>
<td>-7.106 (-28.23)</td>
<td>-7.093 (-24.61)</td>
<td>-7.275 (-38.23)</td>
<td>-7.243 (-26.07)</td>
<td>-7.540 (-28.63)</td>
</tr>
<tr>
<td><strong>Translation Parameters</strong></td>
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<tr>
<td>( \gamma_k )</td>
<td>147.85 (4.10)</td>
<td>138.13 (3.96)</td>
<td>27.87 (9.43)</td>
<td>145.02 (3.89)</td>
<td>100.20 (6.22)</td>
<td>251.99 (8.85)</td>
<td>262.38 (7.07)</td>
<td>42.40 (10.80)</td>
<td>42.02 (3.44)</td>
<td>36.90 (4.51)</td>
<td>111.06 (4.08)</td>
<td>127.26 (4.84)</td>
</tr>
<tr>
<td><strong>Household Sociodemographics</strong></td>
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<tr>
<td>Single person family</td>
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<tr>
<td>Caucasian</td>
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<tr>
<td>Have internet access at home</td>
<td>2.391 (4.51)</td>
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<tr>
<td>Medium annual household income (35-90K)</td>
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<tr>
<td>High annual household income (&gt;90K)</td>
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<tr>
<td>Number of motorized vehicles</td>
<td>-0.122 (-1.85)</td>
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<tr>
<td>Number of bicycles</td>
<td>-0.042 (-1.38)</td>
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<tr>
<td><strong>Individual Sociodemographics and Employment Characteristics</strong></td>
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</tr>
<tr>
<td>Male</td>
<td>-0.310 (-1.72)</td>
<td>0.387 (1.99)</td>
<td>0.680 (2.25)</td>
<td>0.921 (5.12)</td>
<td>0.276 (2.11)</td>
<td>0.544 (3.93)</td>
<td>0.520 (4.15)</td>
<td>0.381 (1.32)</td>
<td>0.161 (1.10)</td>
<td>0.591 (2.68)</td>
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<tr>
<td>Age &lt; 30 yrs</td>
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<tr>
<td>Age &gt; 65 yrs</td>
<td>0.656 (3.31)</td>
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<tr>
<td>Driver’s license</td>
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<tr>
<td>Physically challenged</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td>-0.610 (-1.19)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Employed</td>
<td>-0.390 (-1.60)</td>
<td>-0.414 (-1.79)</td>
<td>-0.177 (-1.58)</td>
<td>0.425 (4.07)</td>
<td></td>
<td></td>
<td></td>
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<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Duration of work/school activity (in minutes)</td>
<td>-0.159 (-3.83)</td>
<td>-0.091 (-1.84)</td>
<td>-0.135 (-6.51)</td>
<td>-0.113 (-3.85)</td>
<td>0.060 (3.02)</td>
<td>-0.118 (-2.46)</td>
<td>-0.093 (-2.88)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
Table 3 (continued). Estimation Results of the Time-use Component of the Joint Model

<table>
<thead>
<tr>
<th>Day of the Week and Seasonal Variables</th>
<th>OH Volunteering</th>
<th>IH Internet use</th>
<th>OH Shopping</th>
<th>IH Socializing</th>
<th>OH Socializing</th>
<th>IH Relaxing</th>
<th>IH Recreation</th>
<th>OH Meals</th>
<th>OH Physically Active Pure Recreation</th>
<th>OH Physically Inactive Pure Recreation</th>
<th>OH Physically Active Recreation</th>
<th>OH Physically Inactive Recreation</th>
</tr>
</thead>
<tbody>
<tr>
<td>Friday</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.382 (3.00)</td>
<td>0.769 (3.01)</td>
<td>0.629 (3.78)</td>
<td>0.322 (2.73)</td>
<td>0.313 (2.33)</td>
<td>0.436 (3.89)</td>
<td>-0.468 (-2.70)</td>
<td>-0.468 (-2.70)</td>
</tr>
<tr>
<td>Summer</td>
<td></td>
<td></td>
<td></td>
<td></td>
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<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fall</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>-0.360 (-3.39)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>-0.401 (-2.33)</td>
<td>-0.401 (-2.33)</td>
</tr>
</tbody>
</table>

Activity-Travel Environment Attributes

| Employment density (#jobs per acre x 10^{-4}) | 0.013 (1.61) | 0.013 (1.61) | 0.013 (1.61) | 0.013 (1.61) |
| #eat-out centers per square mile (x 10^{-7})  |               |               |               | 0.026 (2.65) |
| #physically inactive recreational centers per square mile (x 10^{-7}) |               |               |               |               |
| Presence of more than 4 sports/fitness centers in the residential zone interacted with presence of bicycles in household |               |               |               | 0.187 (2.84) |
| Bicycling facility density (miles of bike lanes per square mile) |               |               |               | 0.116 (2.47) |
| Bicycling facility density interacted with #vehicles/#adults |               |               |               | -0.050 (-1.00) |
Table 4. Impact of Change in ATE Attributes and Sociodemographic Characteristics

<table>
<thead>
<tr>
<th></th>
<th>Predicted % change in aggregate time use in recreational activity purposes</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Household vehicle ownership decreased by 1</td>
</tr>
<tr>
<td>IH Recreation</td>
<td>14.48</td>
</tr>
<tr>
<td>OH Physically Active Pure Recreation</td>
<td>31.47</td>
</tr>
<tr>
<td>OH Physically Inactive Pure Recreation</td>
<td>--</td>
</tr>
<tr>
<td>OH Physically Active Recreation</td>
<td>18.07</td>
</tr>
<tr>
<td>OH Physically Inactive Recreation</td>
<td>--</td>
</tr>
</tbody>
</table>